Robust estimation of the quintile share ratio with bias reduction

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Laeken indicators

- European statistical indicators on poverty and social exclusion
- Lisbon Strategy: coordination of European social policies
- European Statistics on Income and Living Condition (EU-SILC): Coordinatet surveys across European Countries
- Survey methodology not main concern in development of the Laeken indicators: AMELI

Lisbon strategy (2000): "to become the most competitive and dynamic knowledge-based economy in the world capable of sustainable economic growth with more and better jobs and greater social cohesion"

AMELI

- Advanced Methodology for European Laeken Indicators
- ▶ Reserach project of 2008-2011 EU FP7 SSH programme.
- Methodology: robustness, small area estimation, variance estimation, visualisation.
- Development and evaluation by simulation.
- ▶ 10 Partners (4 Universities, 6 NSI)
- April 2008 March 2011
- ► 1.4 M€

Quintile Share Ratio

Inequality of income distribution - Income quintile share ratio

Short description: The ratio of total income received by the 20% of the population with the highest income (top quintile) to that received by the 20% of the population with the lowest income (lowest quintile). Income must be understood as equivalised disposable income.

QSR is much simpler to explain than Gini!

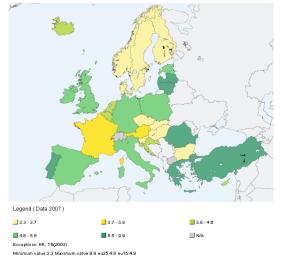
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Official figures from Eurostat

Inequality of income distribution

Income quintile share ratio

 $\mathbf{n}|w$



QSR

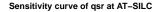
Estimand

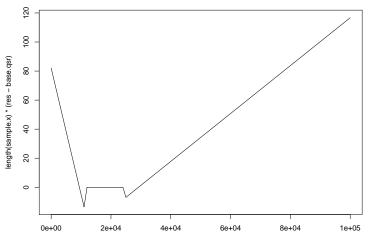
$$\eta(F_U) = \frac{\sum_{i \in U} y_i \, \mathbf{1}\{Q_{0.8}(F_U) \le y_i \le Q_1(F_U)\}}{\sum_{i \in U} y_i \, \mathbf{1}\{Q_0(F_U) \le y_i \le Q_{0.2}(F_U)\}}$$

Estimator: $\hat{\eta} = \eta(F_S)$

- QSR has a breakdown point of 0: Moving a single observation any desired positive value can be produced.
- ► The influence function of the QSR is unbounded above.
- The variability of QSR is high.

 $\mathbf{n} \boldsymbol{w}$





 $\mathbf{n} w$

i.i.d. observations: some Notation

Let $\alpha, \beta \in [0, 1]$, $\alpha < \beta$ and $n\alpha, n\beta \in \{1, \dots, n\}$ Trimmed mean

$$T_{\alpha;\beta}(F_n) := \frac{1}{n(\beta - \alpha)} \sum_{i=n\alpha+1}^{n\beta} X_{(i)}$$

The trimmed mean is an L-functional:

$$T_{\alpha;\beta}(F) = \int_0^1 F^{-1}(y) \frac{\mathbf{1}\{\alpha \le y \le \beta\}}{\beta - \alpha} dy$$

Robustification

- Parametric modelling and classical robustification (e.g. M-estimation)
- Parametric modeling of the tail above a threshold
- Trimming of extreme observations (Cowell and Victoria-Feser)
- AMELI: Nonparametric Robustification of QSR under complex sampling (Hulliger and Schoch 2009)

Trimming of extremes

Extreme observations are (symmetrically) trimmed from both tails. Upper (α_u) and lower (α_l) trimming proportions are equal.

$$\hat{\eta}_{TQSR}(\alpha_I, \alpha_u) = \frac{T_{0.8; 1-\alpha_u}(F)}{T_{\alpha_I; 0.2}(F)}$$

Both trimmings move TQSR down: High Bias!

Bias compensation

Trimming both quintiles above

$$\hat{\eta}_{BQSR}(\alpha_l, \alpha_u) = \frac{T_{0.8; 1-\alpha_u}(F)}{T_{0; 0.2-\alpha_l}(F)}$$

The trimming of numerator and denominator compensate at least partialley.

For a given (small) upper trimming proportion α_u there is a α_l such that BQSR is unbiased:

$$\hat{\eta}_{BQSR}\left(\alpha_{I}(\alpha_{u}),\alpha_{u}\right)=\hat{\eta}$$

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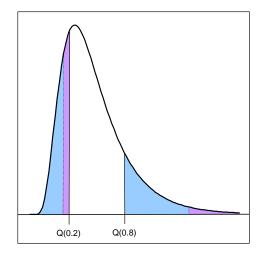
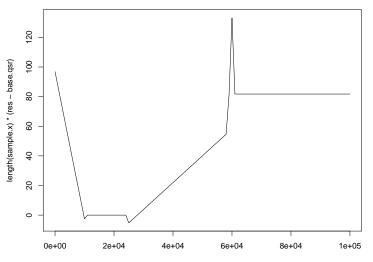


Figure: Idealized density function

 $\mathbf{n} \boldsymbol{w}$





Correction term

- Trim only numerator of QSR: $T_{0.8; 1-\alpha_u}(F)$
- Use approximate correction term T_{0.8; 1-α_u(F)} by considering "double trimming".

$$\hat{\eta}_{CQSR}(\alpha_u) = \frac{T_{0.8; \ 1-\alpha_u}(F)}{T_{0; \ 0.2}(F)} \cdot \frac{T_{0.8; \ 1-\alpha_u}(F)}{T_{0.8; \ 1-2\alpha_u}(F)}$$

Sampling adaptations

- The functionals in T_{α;β}(F) have to be estimated from the sample, taking the sample design into account.
- Plug in an estimate of $F_U(t)$: $F_S(t) = \sum_S w_i \mathbf{1}\{x_i \le t\} / \sum_S w_i$:

$$\hat{T} = T(F_S)$$

Weighted trimmed means are straightforward
⇒ Shao, Annals of Statistics 1994

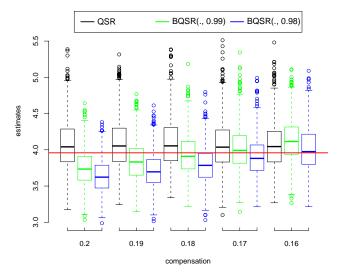
Variance estimation under sampling

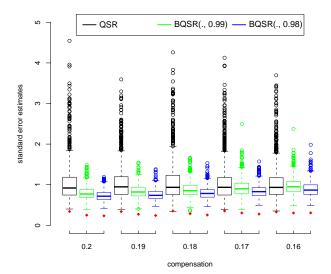
- 1. Estimate the variance covariance matrix of the influence function of the involved estimators taking into account the sample design.
- 2. Apply the Delta method.
- If T_j(F_n) = ∑ⁿ_i w_ix_{ij} linearize first and take directional derivaitive (Woodruff 1971)(Anderson and Nordberg 1994)(Münnich 2008)

Some results with AT-SILC and BQSR

- Public Use Sample of AT-SILC 2004 (Sample for Research)
- Sample design: Stratified sample of households, proportional allocation. Rotational design. Complex non-response adjustment.
- Unit: persons (n = 4626, p = 240)
- Simulation with simple-random samples of size n = 400 and r = 1000 replicates.

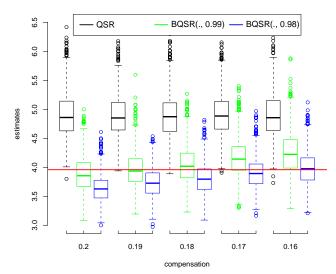




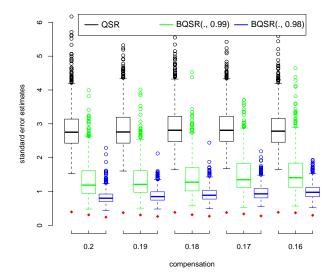


Estimates of SE[BQSR], N=400

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Estimates of BQSR, N=400, Outliers=c(1.5e5, 2e5, 2e5, 2.5e5)



Estimates of SE[BQSR], N=400, Outliers=c(1.5e5, 2e5, 2e5, 2.5e5)

Final remarks

- Work in progress!
- ► Variance estimator to be enhanced.
- \blacktriangleright We do not change the parameter η to estimate but the estimator.
- Variance estimation will be extensively tested in AMELI-simulations.
- > The problem of choosing a tuning constant remains.

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http://ameli.surveystatistics.net
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