

Handling Missing Data for Indicators

Susanne Rässler

**Institute for Employment Research & Federal Employment Agency
Nürnberg, Germany**

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Agenda

- **Missing data**
- **Multiple imputation principle**
- **Indicators are missing**
- **Multivariate MI approach**
- **Alternative approaches via flexible chained equations**
- **Conclusions**

Missing Data - “everybody has them, nobody wants them”

Unit no.	Gender	Age	Education	Health state	Personal Net-Income	...
1	female	40-45	high	good	?	...
2	male	30-35	middle	poor	4500-5000	...
3	female	>60	?	poor	4000-4500	...
4	male	20-25	high	?	?	...
5	male	20-25	low	?	1500-2000	...
6	female	30-35	low	good	1500-2000	...
...

Case Deletion

Unit no.	Gender	Age	Education	Health state	Personal Net-Income	...
2	male	30-35	middle	poor	4500-5000	...
6	female	30-35	low	good	1500-2000	...
...

Missingness may be either

- MCAR (missing completely at random),
- MAR (missing at random), or
- MNAR (missing not at random)

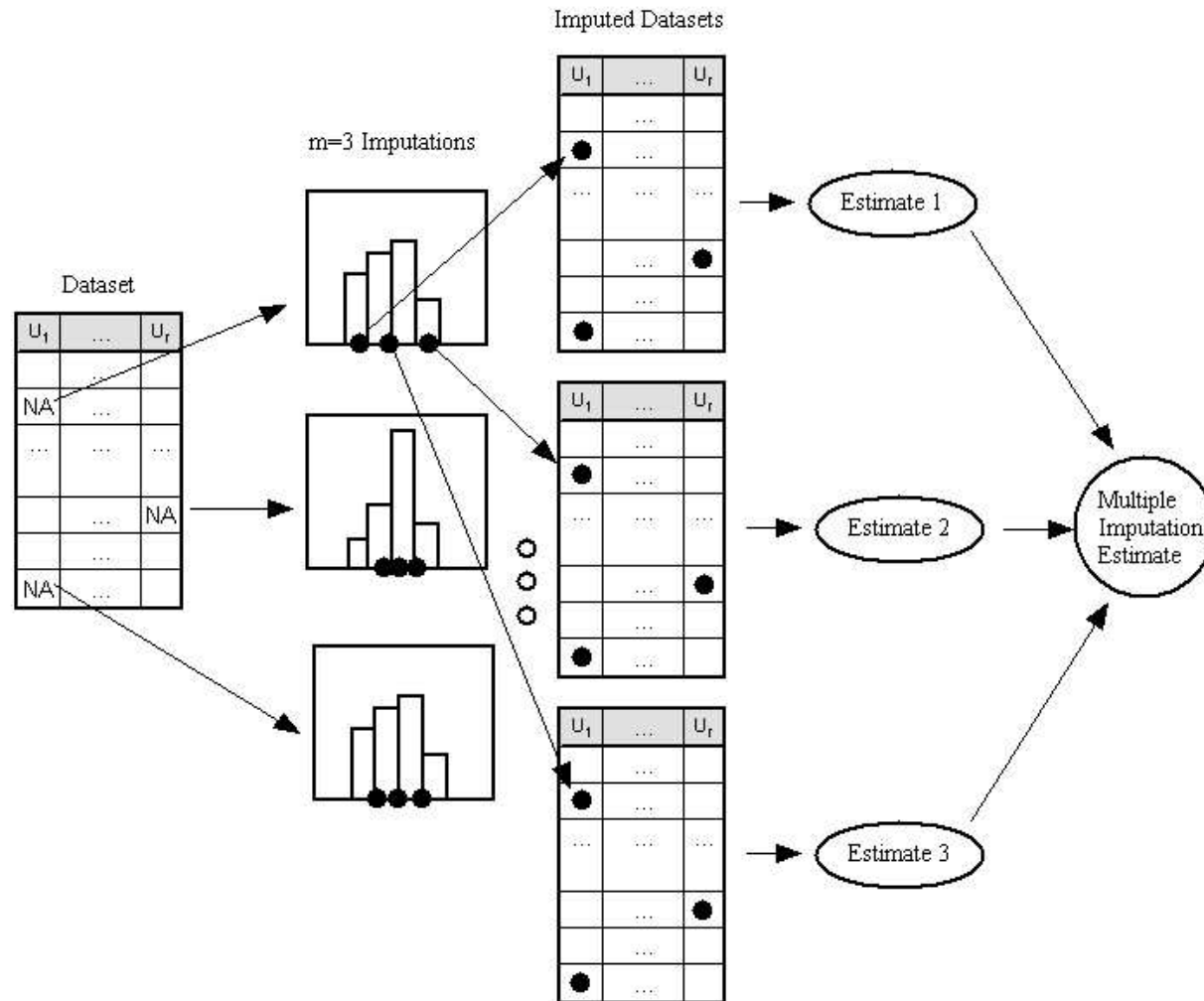
(Rubin and Little 1987, 2002)

⇒ In multivariate analysis often 30% to 40% of the data are lost with case deletion assuming MCAR!

Handling missing data

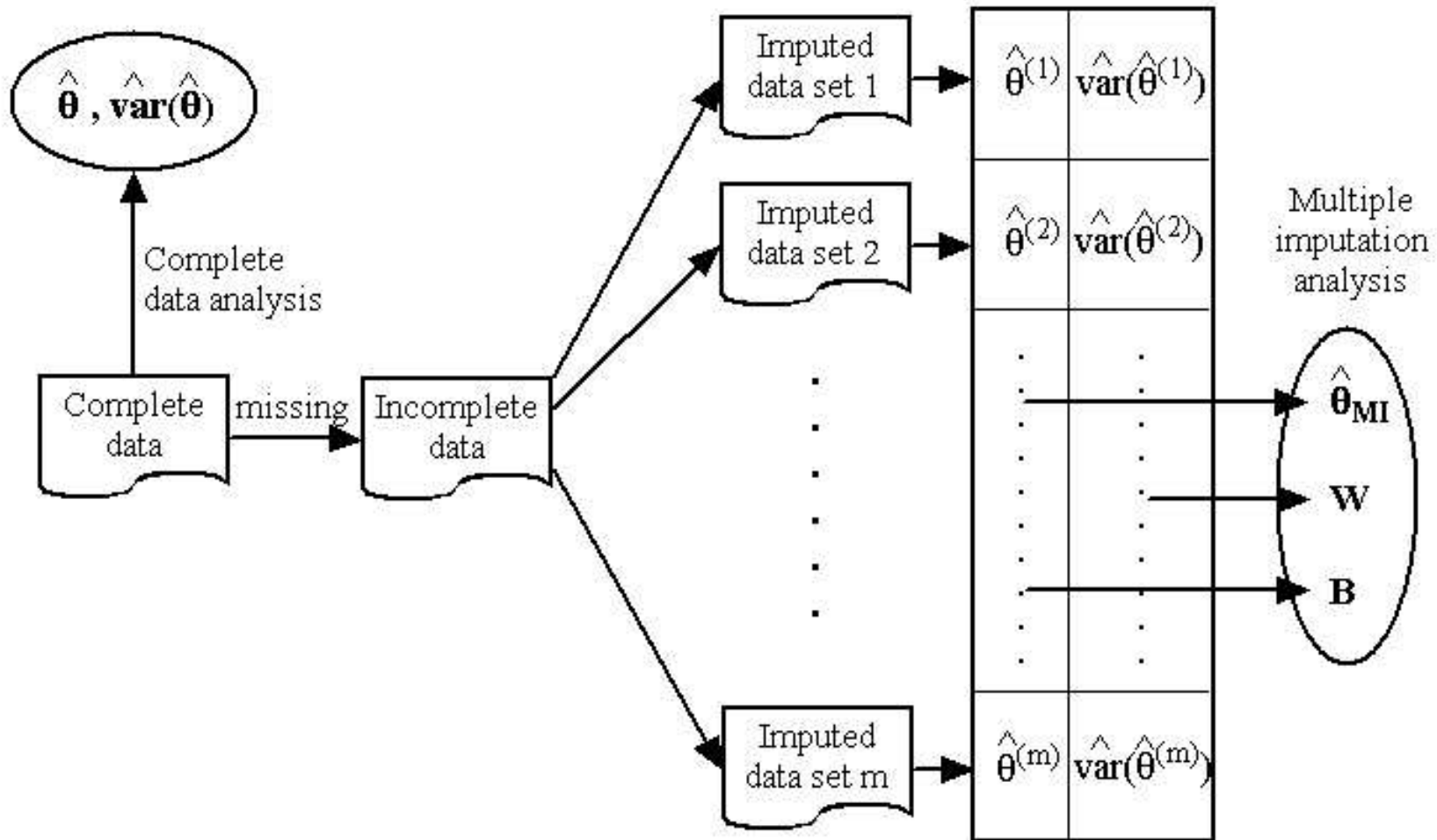
- Procedures based on the available cases only, i.e., only those cases that are completely recorded for the variables of interest
 - Weighting procedures such as Horvitz-Thompson type estimators or raking estimators that adjust for nonresponse
 - Single imputation and correction of the variance estimates to account for imputation uncertainty
 - Multiple imputation (MI) according to Rubin (1978, 1987) and standard complete-case analysis
 - Model-based corrections of parameter estimates such as the expectation-maximization (EM) algorithm
- ⇒ We regard multiple imputation as most flexible for multipurpose complex surveys such as KEI

The Multiple Imputation Principle (1)



⇒ MI reflects **uncertainty about which value to impute**

The Multiple Imputation Principle (2)



⇒ Correct MI analysis is based on an **analysis of variance**

The Multiple Imputation Principle (3)

- **Estimates:** with complete data, tests and intervals based on the normal approximation should be appropriate (Rubin 1978, 1987, or t -approximation, Barnard and Rubin 1999); i.e.,

$$(\hat{\theta} - \theta) / \sqrt{\widehat{\text{var}}(\hat{\theta})} \sim N(0, 1)$$

- **Produce m completed data sets and calculate $\hat{\theta}^{(j)}$ and $\widehat{\text{var}}(\hat{\theta}^{(j)})$, $j = 1, 2, \dots, m$**
- **Multiple imputation estimate $\hat{\theta}_{MI} = \frac{1}{m} \sum_{j=1}^m \hat{\theta}^{(j)}$**
- **Its estimated total variance is $T = W + (1 + \frac{1}{m})B$ with**
“within-imputation” variance $W = \frac{1}{m} \sum_{j=1}^m \widehat{\text{var}}(\hat{\theta}^{(j)})$ and
“between-imputation” variance $B = \frac{1}{m-1} \sum_{j=1}^m (\hat{\theta}^{(j)} - \hat{\theta}_{MI})^2$

⇒ **Tests can be based on $(\hat{\theta}_{MI} - \theta) / \sqrt{T} \sim t_v$ with $v = (m - 1) \left(1 + \frac{W}{(1+m^{-1})B}\right)^2$**

Basic Principle of Multiple Imputation Procedures

- Create m independent random draws of the missing data according to their posterior predictive distribution

$$f_{Y_{mis}|Y_{obs}}(y_{mis}|y_{obs}) = \int f_{Y_{mis}|Y_{obs},\Theta}(y_{mis}|y_{obs},\theta) f_{\Theta|Y_{obs}}(\theta|y_{obs}) d\theta$$

- **Realization either by**
 - (1) random draws of the parameters Θ according to their **observed-data posterior distribution** $f_{\Theta|Y_{obs}}$ as well as
 - (2) random draws of Y_{mis} according to their conditional predictive distribution $f_{Y_{mis}|Y_{obs},\Theta}$ for actual draws of Θ .
- **or realization iteratively (MCMC, data augmentation) by**
 - (1) random draws of the parameters Θ according to their **complete-data posterior distribution** $f_{\Theta|Y_{obs},Y_{mis}}$ for actual draws of Y_{mis} as well as
 - (2) random draws of Y_{mis} according to their conditional predictive distribution $f_{Y_{mis}|Y_{obs},\Theta}$ for actual draws of Θ .

Indicators are missing

- **Countries:** EU + USA + Japan
 - **Time period:** 1995 ... 2002/03, early *estimates* for 2003/04
 - **Indicators:**
 - GERD** Gross domestic expenditure for R & D per capita (POP)
 - PhD** Total new science and technology PhDs per capita
 - FTE** Total researchers (FTE) per capita
 - GFCF** Total gross fixed capital formation (excl. building) per capita
 - EGov** E-government
 - TEE** Total education expenditure per capita
 - LLL** Life-long learning (per population aged 25-64 years participating in education and training; POP1)
- ⇒ **Some indicators are missing at the most recent point of time**

Where to Go Intermediate: The multivariate model for KEI

- Data augmentation algorithm using the multivariate linear mixed-effects model (Schafer & Yucel 2002)

$$Y_i = X_i\beta + Z_ib_i + \epsilon_i, \quad i = 1, 2, \dots, n$$

Y_i	=	$(T \times r)$ matrix of indicators
X_i	=	$(T \times p)$ matrix of covariates
Z_i	=	$(T \times q)$ matrix of covariates with Z_i basically $\in X_i$
β	=	$(p \times r)$ matrix of fixed effects
$\text{vec}(b_i)$	\sim	$N_{qr}(0, \Psi)$ vector of random effects
$\text{vec}(\epsilon_i)$	\sim	$N_{Tr}(0, \Sigma \otimes I_T)$ vector of random errors
Ψ^{-1}	\sim	$Wishart_{qr}(a, B)$, a, B hyperparameter
Σ^{-1}	\sim	$Wishart_r(c, D)$, c, D hyperparameter

- times of measurement t incorporated into X_i and possibly Z_i .
- allows unequal spacing, time-varying covariates, unbalanced panels for T_i , correlation between indicators.

Univariate Multiple Imputation Models for Complex Data

Simple case with 3 variables A , B and C each with missing data (Rubin 2003, applied in the NMES):

- “Begin by arbitrarily filling in all missing B and C .
- Fit a model of $A|B, C$ using those units where A is observed and impute the missing A values.
- Toss the imputed B values and fit a model of $B|A, C$ using those units where B is observed and impute the missing B values.
- Toss the imputed C values and fit a model of $C|A, B$ using those units where C is observed and impute the missing C values.
- Iterate...”

⇒ Great flexibility due to the possible conditional specifications!

Univariate KEI-Imputations Based on PAN (1)

- Assume indicators are missing at random (MAR)
- Fit univariate mixed-effects model for each KEI indicator separately (SPLUS library pan by Schafer 1997):

$$y_i = X_i\beta + Z_ib_i + \epsilon_i, \quad i = 1, 2, \dots, n$$

y_i = $(y_{i1}, y_{i2}, \dots, y_{iT})$ **KEI indicator**

X_i = **(intercept, time)**

Z_i = **(intercept)**

β_0, β_1 = **fixed effects of intercept and time**

b_i $\sim N(0, \psi)$ **random effect for country i**

ϵ_i $\sim N_T(0, \sigma^2 I_T)$ **random errors**

- Leads to model $y_i \sim N_T(X_i\beta, \psi + \sigma^2 I_T)$ for $i = 1, 2, \dots, n$ such that

$$\text{Cov}(y_{it}, y_{js}) = \begin{cases} \psi, & t, s = 1, 2, \dots, T, t \neq s, i = j \\ \psi + \sigma^2, & t, s = 1, 2, \dots, T, t = s, i = j \\ 0, & \text{else, i.e. for all } i \neq j \end{cases}$$

First Results of KEI-Imputations Based on PAN (2)

- Generate $m = 10$ imputations after a burn-in period of 1000 Gibbs cycles.
- ACF's of ψ , σ^2 and β suggest quick convergence
- Lags of 100 between each imputation are used
- To Do:
 - allow correlation between indicators \Rightarrow Pan for KEI according to Schafer & Yucel (2002)
 - allow for heteroscedasticity \Rightarrow possibly with approach Schafer & Yucel (2002)
 - allow for flexible serial autocorrelation \Rightarrow future research
 - allow for spacial autocorrelation \Rightarrow future research

Conclusions

- MI is in general applicable when the complete-data estimates are asymptotically normal (like ML estimates are) or t distributed.
 - The regression switching approach seems to be quite promising in large data sets and for high amounts of missing values.
 - Even in the context of “mass imputation”, such as split questionnaire survey designs and data fusion we find good frequentist properties.
 - In the U.S. applied for MI in the NHANES (split project) and NMES.
 - The basic routines are already implemented in MICE (SPLUS and R version) and IVEware, Raghunathan’s SAS callable application.
- ⇒ Multiple imputation displays nonresponse uncertainty while using standard complete-case analysis!

References

- **Barnard, J., Rubin, D.B. (1999), Small-Sample Degrees of Freedom with Multiple Imputation, *Biometrika*, 86, 948-955.**
- **Rubin, D.B. (1978), Multiple Imputations in Sample Surveys - A Phenomenological Bayesian Approach to Nonresponse, (with discussion and reply), *Proceedings of the Survey Research Methods Section of the American Statistical Association*, 20-34.**
- **Rubin, D.B. (1987), *Multiple Imputation for Nonresponse in Surveys*, Wiley, New York.**
- **Little, R.J.A., Rubin, D.B. (1987, 2002), *Statistical Analysis with Missing Data*, Wiley, New York.**
- **Schafer, J.L. (1997), *Analysis of Incomplete Multivariate Data*, Chapman and Hall, London.**
- **Schafer, J.L., Yucel, R. (2002) Computational Strategies for Multivariate Linear Mixed-Effects Models with Missing Values, *Journal of Computational and Graphical Statistics*, 11, 437-457.**